

# Bayesian Learning

Lecture 2 - Poisson data. Prior elicitation. Invariant priors.

**Mattias Villani**

Department of Statistics  
Stockholm University



mattiasvillani.com



@matvil



@matvil



mattiasvillani

# Lecture overview

- The **Poisson model**
- **Summarizing a posterior distribution**
- **Conjugate priors**
- **Prior elicitation**
- **Jeffreys' prior**

# Poisson model

## ■ Model

$$y_1, \dots, y_n | \theta \stackrel{iid}{\sim} \text{Pois}(\theta)$$

## ■ Poisson distribution

$$p(y) = \frac{\theta^y e^{-\theta}}{y!}$$

## ■ Likelihood from iid Poisson sample $y = (y_1, \dots, y_n)$

$$p(y|\theta) = \left[ \prod_{i=1}^n p(y_i|\theta) \right] \propto \theta^{(\sum_{i=1}^n y_i)} \exp(-\theta n),$$

## ■ Prior

$$p(\theta) \propto \theta^{\alpha-1} \exp(-\theta\beta) \propto \text{Gamma}(\alpha, \beta)$$

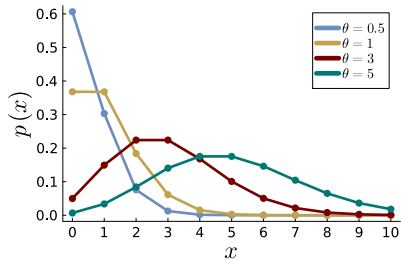
# Poisson distribution

$X \sim \text{Pois}(\theta)$   
for  $X \in 0, 1, 2, \dots$

$$p(x) = \frac{\theta^x e^{-\theta}}{x!}$$

$$\mathbb{E}(X) = \theta$$

$$\mathbb{V}(X) = \theta$$



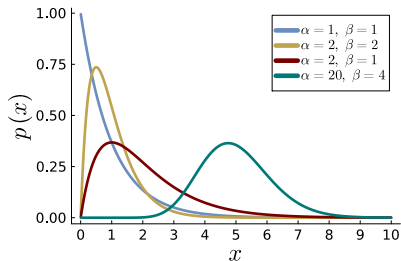
# Gamma distribution

$$X \sim \text{Gamma}(\alpha, \beta)$$

$$p(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$$

$$\mathbb{E}(X) = \frac{\alpha}{\beta}$$

$$\mathbb{V}(X) = \frac{\alpha}{\beta^2}$$



# Poisson posterior

## ■ Posterior

$$\begin{aligned} p(\theta|y_1, \dots, y_n) &\propto \left[ \prod_{i=1}^n p(y_i|\theta) \right] p(\theta) \\ &\propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n) \theta^{\alpha-1} \exp(-\theta\beta) \\ &= \theta^{\alpha + \sum_{i=1}^n y_i - 1} \exp[-\theta(\beta + n)], \end{aligned}$$

which is proportional to  $\text{Gamma}(\alpha + \sum_{i=1}^n y_i, \beta + n)$ .

## ■ Bayesian updating

### Conjugate analysis - Poisson model

**Model:**  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\theta)$

**Prior:**  $\theta \sim \text{Gamma}(\alpha, \beta)$

**Posterior:**  $\theta|x_1, \dots, x_n \sim \text{Gamma}(\alpha + \sum_{i=1}^n x_i, \beta + n)$

# Example - Number of bids in eBay auctions

## ■ Data:

- ▶ Number of placed bids in  $n = 1000$  eBay coin auctions.
- ▶ Sum of counts:  $\sum_{i=1}^n y_i = 3635$ .
- ▶ Average number bids per auction:  $\bar{y} = 3635/1000 = 3.635$ .

## ■ Prior: $\alpha = 2$ , $\beta = 1/2$ .

$$\mathbb{E}(\theta) = \frac{\alpha}{\beta} = 4$$

$$\mathbb{S}(\theta) = \sqrt{\frac{\alpha}{\beta^2}} \approx 2.823$$

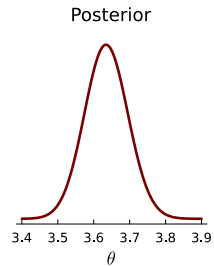
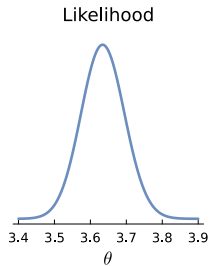
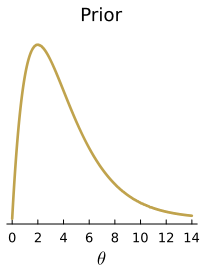
## ■ Posterior mean and posterior standard deviation

$$\theta|\mathbf{y} \sim \text{Gamma}(\alpha + \sum_{i=1}^n y_i, \beta + n) = \text{Gamma}(3637, 1000.5)$$

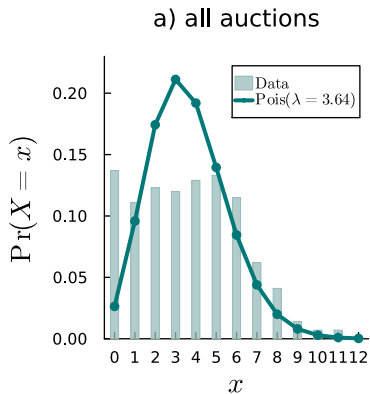
$$\mathbb{E}(\theta|\mathbf{y}) = \frac{\alpha + \sum_{i=1}^n y_i}{\beta + n} = \frac{2 + 3635}{1/2 + 1000} \approx 3.635$$

$$\mathbb{S}(\theta|\mathbf{y}) = \left( \frac{\alpha + \sum_{i=1}^n y_i}{(\beta + n)^2} \right)^{1/2} \approx 0.060$$

# eBay data - Posterior of $\theta$



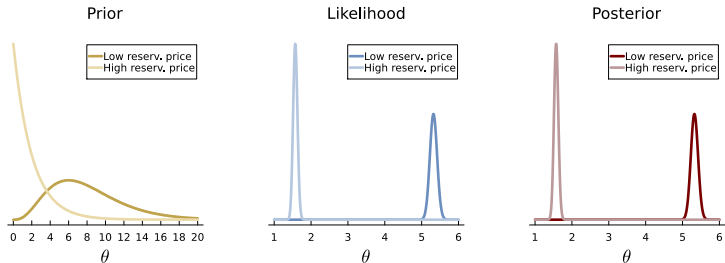
# eBay data - model fit at $\theta = \mathbb{E}(\theta|\mathbf{x})$



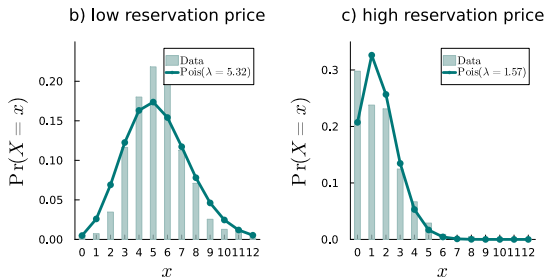
# eBay - low/high seller's reservation price

- The data is very heterogenous. Some auctions start with very high reservations prices (lowest price accepted by the seller).
- Split the data into auctions with low/high reservation prices.
- **Low reservation price auctions:**
  - ▶  $n = 550$  eBay coin auctions.
  - ▶ Posterior mean: 5.321 bids.
- **High reservation price auctions:**
  - ▶  $n = 450$  eBay coin auctions.
  - ▶ Posterior mean: 1.576 bids.

# eBay data split on reservation price



# eBay data - model fit at $\mathbb{E}(\theta|\mathbf{x})$



- Better fits, but still not good enough.
- Even better: **Poisson regression** with reservation price as continuous covariate.

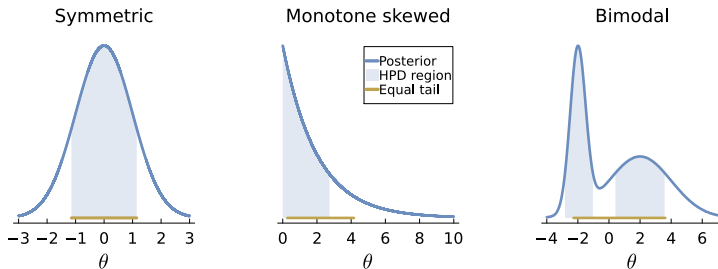
# Posterior intervals

- **Bayesian 95% credible interval**: the unknown parameter  $\theta$  lies in the interval with 0.95 probability **conditional on the observed data**.
- 95% **equal-tail interval**: from 2.5% to 97.5% percentile.
- Approximate 95% **credible interval**

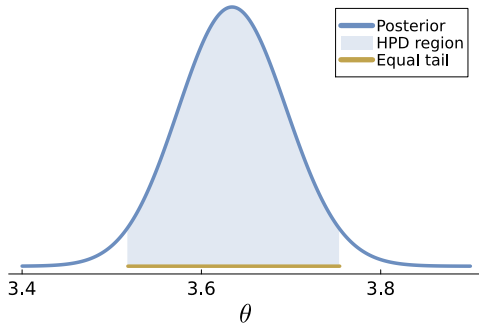
$$\mathbb{E}(\theta|\mathbf{y}) \pm 1.96 \cdot \mathbb{S}(\theta|\mathbf{y})$$

- **Highest Posterior Density (HPD)** interval contains the  $\theta$  values with highest posterior pdf.

# Illustration of different interval types



# Credible intervals - eBay auction data



# Conjugate priors

- Normal likelihood: Normal prior  $\rightarrow$  Normal posterior.
- Bernoulli likelihood: Beta prior  $\rightarrow$  Beta posterior.
- Poisson likelihood: Gamma prior  $\rightarrow$  Gamma posterior.
- **Conjugate priors**: A prior is conjugate to a model if the prior and posterior belong to the **same distributional family**.

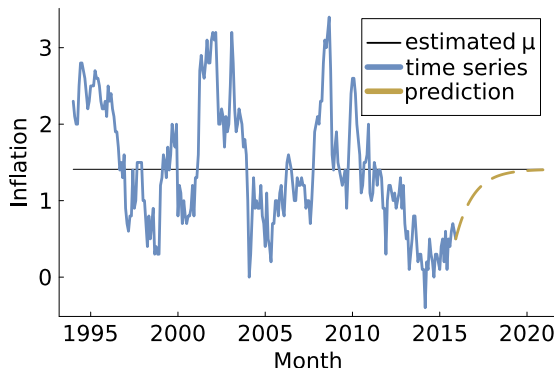
# Autoregressive time series model

- **Autoregressive process** of order  $p$  -  $AR(p)$

$$y_t = \mu + \phi_1(y_{t-1} - \mu) + \dots + \phi_p(y_{t-p} - \mu) + \varepsilon_t, \quad \varepsilon_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

- Unconditional mean:  $\mathbb{E}(y_t) = \mu$ . Long run forecast attraction.

$$\mathbb{E}(y_{T+h}|y_{1:T}) \rightarrow \mu \text{ as } h \rightarrow \infty.$$



# Prior elicitation - AR(p)

## ■ Autoregressive process

$$y_t = \mu + \phi_1(y_{t-1} - \mu) + \dots + \phi_p(y_{t-p} - \mu) + \varepsilon_t$$

■ **Expert prior** on the unconditional mean:  $\mu \sim N(\mu_0, \tau_0^2)$ .

■ **Regularization prior** on  $\phi_1, \dots, \phi_p$

$$\phi_k \sim N\left(\mu_k, \sigma^2 \frac{\tau^2}{k^2}\right) \text{ independently apriori}$$

- ▶ Prior mean on persistent AR(1):  $\mu_1 = 0.8, \mu_2 = \dots = \mu_p = 0$
- ▶  $\mathbb{V}(\phi_k) = \sigma^2 \frac{\tau^2}{k^2}$ . Coeff on “longer” lags more likely to be small.

## ■ Hierarchical prior

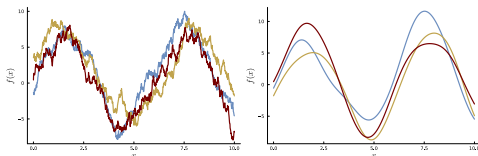
- ▶ Hard to specify  $\tau^2$ ? Put a prior on it!
- ▶  $\phi_k | \tau^2 \sim N\left(\mu_k, \sigma^2 \frac{\tau^2}{k^2}\right)$  and  $\tau^2 \sim \chi_\nu^2$ .
- ▶ Gives a posterior on global shrinkage  $\tau^2$ .

# Prior elicitation

## Smoothness priors

- ▶ a version of regularization priors
- ▶ nonlinear regression function  $f(\cdot)$  is believed to be smooth

$$y = f(x) + \varepsilon$$



## Noninformative priors

- ▶ **Uniform:**  $\theta \sim \text{Beta}(1, 1)$ .  
Issue 1: same as prior sample with one success and one failure.  
Issue 2: not uniform for  $\phi = \log \frac{\theta}{1-\theta}$ .
- ▶ **Zero prior sample size:**  $\theta \sim \text{Beta}(\epsilon, \epsilon)$  with  $\epsilon \downarrow 0$ .  
Posterior  $\rightarrow \text{Beta}(s, f)$ .  
Issue: posterior is improper if  $s = 0$  or  $f = 0$ .

# Invariant prior

## ■ Observed information

$$J_{\mathbf{x}}(\hat{\theta}) = - \frac{d^2 \ln p(\mathbf{x}|\theta)}{d\theta^2} \Big|_{\theta=\hat{\theta}}$$

## ■ Fisher information

$$I(\theta) = E_{\mathbf{x}|\theta} (J_{\mathbf{x}}(\theta))$$

## ■ Jeffreys' rule to construct prior

$$p(\theta) = I(\theta)^{1/2}.$$

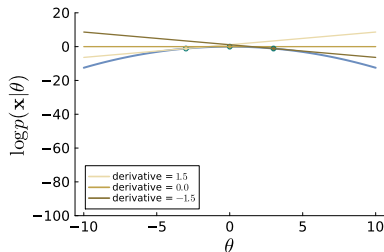
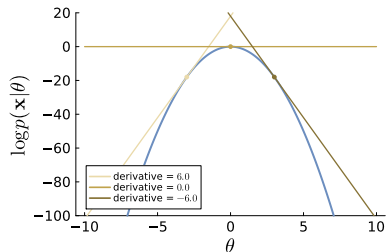
## ■ Invariance under 1:1 parameter transformation $\phi = g(\theta)$ .

Example:  $\phi = \log \frac{\theta}{1-\theta}$ .

▶ Specify  $p_{\theta}(\theta)$  directly

▶ Specify  $p_{\phi}(\phi)$  and then obtain  $p_{\theta}(\theta) = p_{\phi}(g^{-1}(\theta)) \left| \frac{dg^{-1}(\theta)}{d\theta} \right|$ .

# Second derivative measures curvature



# Jeffreys' prior for negative binomial sampling

- Jeffreys' prior:

$$n|\theta \stackrel{iid}{\sim} \text{NegBin}(s, \theta).$$

$$\ln p(\mathbf{x}|\theta) = \ln \binom{n-1}{s-1} + s \ln \theta + f \ln(1 - \theta)$$

$$\frac{d^2 \ln p(\mathbf{x}|\theta)}{d\theta^2} = -\frac{s}{\theta^2} - \frac{f}{(1 - \theta)^2}$$

$$I(\theta) = \frac{s}{\theta^2} + \frac{E_{n|\theta}(n - s)}{(1 - \theta)^2} = \frac{s}{\theta^2} + \frac{s/\theta - s}{(1 - \theta)^2} = \frac{s}{\theta^2(1 - \theta)}$$

- Thus, the Jeffreys' prior is

$$p(\theta) = |I(\theta)|^{1/2} \propto \theta^{-1}(1 - \theta)^{-1/2} \propto \text{Beta}(\theta|0, 1/2).$$

- Jeffreys' prior is **improper**, but the posterior is proper:  
 $\theta|n \sim \text{Beta}(s, f + 1/2)$  which is proper since  $s \geq 1$ .
- Jeffreys' prior **violates the likelihood principle** because  $I(\theta)$  is sampling-based.